

Regression with a Binary Dependent Variable, Part II

Announcements: No class Mon 10/8; PS4 due Tue 10/9, 9am

Outline

1. Probit with multiple regressors
2. Logit
3. Logit and probit example: *HMDA data*
4. Maximum likelihood estimation
5. Ordered probit (ordered categorical data)

Probit with multiple regressors

```
. probit deny p_irat black, r
```

```
Iteration 0:   log likelihood = -872.0853  
Iteration 1:   log likelihood = -800.88504  
Iteration 2:   log likelihood = -797.1478  
Iteration 3:   log likelihood = -797.13604
```

Probit estimates

```
Number of obs   =          2380  
Wald chi2(2)    =          118.18  
Prob > chi2     =           0.0000  
Pseudo R2      =           0.0859
```

```
Log likelihood = -797.13604
```

		Robust				[95% Conf. Interval]	
deny		Coef.	Std. Err.	z	P> z		
p_irat		2.741637	.4441633	6.17	0.000	1.871092	3.612181
black		.7081579	.0831877	8.51	0.000	.545113	.8712028
_cons		-2.258738	.1588168	-14.22	0.000	-2.570013	-1.947463

We'll go through the estimation details later...

STATA Example, ctd.: predicted probit probabilities

```
. probit deny p_irat black, r
```

Probit estimates

```
Number of obs   =      2380
Wald chi2(2)    =      118.18
Prob > chi2     =      0.0000
Pseudo R2      =      0.0859
```

```
Log likelihood = -797.13604
```

```
-----+-----
            |               Robust
            |               Coef.   Std. Err.      z    P>|z|     [95% Conf. Interval]
-----+-----
p_irat     |    2.741637   .4441633     6.17   0.000   1.871092   3.612181
black      |    .7081579   .0831877     8.51   0.000   .545113   .8712028
_cons     |   -2.258738   .1588168    -14.22  0.000  -2.570013  -1.947463
-----+-----
```

```
. scalar z1 = _b[_cons]+_b[p_irat]*.3+_b[black]*0
```

```
. display "Pred prob, p_irat=.3, white: " normprob(z1)
```

```
Pred prob, p_irat=.3, white: .07546603
```

NOTE: `_b[_cons]` is the estimated intercept (-2.258738)
`_b[p_irat]` is the coefficient on `p_irat` (2.741637)
`scalar` creates a new scalar which is the result of a calculation
`display` prints the indicated information to the screen
`normprob(z1)` computes the cumulative normal probability $\leq z1$

STATA Example, ctd.

$$\begin{aligned} & \Pr(\textit{deny} = 1 \mid P / I, \textit{black}) \\ & = \Phi(-2.26 + 2.74 \times P/I \textit{ ratio} + .71 \times \textit{black}) \\ & \qquad \qquad \qquad (.16) \quad (.44) \qquad \qquad \qquad (.08) \end{aligned}$$

- Is the coefficient on *black* statistically significant?
- Estimated effect of race for *P/I ratio* = .3:

$$\Pr(\textit{deny} = 1 \mid .3, 1) = \Phi(-2.26 + 2.74 \times .3 + .71 \times 1) = \Phi(-0.73) = .233$$

$$\Pr(\textit{deny} = 1 \mid .3, 0) = \Phi(-2.26 + 2.74 \times .3 + .71 \times 0) = \Phi(-1.44) = .075$$

- Difference in rejection probabilities = .158 (15.8 percentage points)
- *Still plenty of room still for omitted variable bias!*

STATA Example: HMDA data – Logit regression

```
. logit deny p_irat black, r;
```

```
Iteration 0:    log likelihood =  -872.0853        Later...
```

```
Iteration 1:    log likelihood =  -806.3571...
```

Logit estimates

Number of obs = 2380

Wald chi2(2) = 117.75

Prob > chi2 = 0.0000

Pseudo R2 = 0.0876

Log likelihood = -795.69521

		Robust					
deny	Coef.	Std. Err.	z	P> z	[95% Conf. Interval]		
p_irat	5.370362	.9633435	5.57	0.000	3.482244	7.258481	
black	1.272782	.1460986	8.71	0.000	.9864339	1.55913	
_cons	-4.125558	.345825	-11.93	0.000	-4.803362	-3.447753	

```
. dis "Pred prob, p_irat=.3, white: "
```

```
> 1/(1+exp(-(_b[_cons]+_b[p_irat]*.3+_b[black]*0)));
```

```
Pred prob, p_irat=.3, white: .07485143
```

```
NOTE: the probit predicted probability is .07546603
```

Predicted probabilities from estimated probit and logit models usually are very close.

The loan officer's decision

- Loan officer uses key financial variables:
 - *P/I ratio*
 - housing expense-to-income ratio
 - loan-to-value ratio
 - personal credit history
- The decision rule is nonlinear:
 - loan-to-value ratio $> 80\%$
 - loan-to-value ratio $> 95\%$
 - credit score
- Illegal to use “protected class” information (gender, race...)

TABLE 11.1 Variables Included in Regression Models of Mortgage Decisions

Variable	Definition	Sample Average
<i>Financial Variables</i>		
<i>P/I ratio</i>	Ratio of total monthly debt payments to total monthly income	0.331
<i>housing expense-to-income ratio</i>	Ratio of monthly housing expenses to total monthly income	0.255
<i>loan-to-value ratio</i>	Ratio of size of loan to assessed value of property	0.738
<i>consumer credit score</i>	1 if no “slow” payments or delinquencies 2 if one or two slow payments or delinquencies 3 if more than two slow payments 4 if insufficient credit history for determination 5 if delinquent credit history with payments 60 days overdue 6 if delinquent credit history with payments 90 days overdue	2.1
<i>mortgage credit score</i>	1 if no late mortgage payments 2 if no mortgage payment history 3 if one or two late mortgage payments 4 if more than two late mortgage payments	1.7
<i>public bad credit record</i>	1 if any public record of credit problems (bankruptcy, charge-offs, collection actions) 0 otherwise	0.074

Additional Applicant Characteristics

<i>denied mortgage insurance</i>	1 if applicant applied for mortgage insurance and was denied, 0 otherwise	0.020
<i>self-employed</i>	1 if self-employed, 0 otherwise	0.116
<i>single</i>	1 if applicant reported being single, 0 otherwise	0.393
<i>high school diploma</i>	1 if applicant graduated from high school, 0 otherwise	0.984
<i>unemployment rate</i>	1989 Massachusetts unemployment rate in the applicant's industry	3.8
<i>condominium</i>	1 if unit is a condominium, 0 otherwise	0.288
<i>black</i>	1 if applicant is black, 0 if white	0.142
<i>deny</i>	1 if mortgage application denied, 0 otherwise	0.120

TABLE 11.2 Mortgage Denial Regressions Using the Boston HMDA DataDependent variable: *deny* = 1 if mortgage application is denied, = 0 if accepted; 2380 observations.

Regression Model Regressor	LPM (1)	Logit (2)	Probit (3)	Probit (4)	Probit (5)	Probit (6)
<i>black</i>	0.084** (0.023)	0.688** (0.182)	0.389** (0.098)	0.371** (0.099)	0.363** (0.100)	0.246 (0.448)
<i>P/I ratio</i>	0.449** (0.114)	4.76** (1.33)	2.44** (0.61)	2.46** (0.60)	2.62** (0.61)	2.57** (0.66)
<i>housing expense-to-income ratio</i>	-0.048 (.110)	-0.11 (1.29)	-0.18 (0.68)	-0.30 (0.68)	-0.50 (0.70)	-0.54 (0.74)
<i>medium loan-to-value ratio</i> ($0.80 \leq \text{loan-value ratio} \leq 0.95$)	0.031* (0.013)	0.46** (0.16)	0.21** (0.08)	0.22** (0.08)	0.22** (0.08)	0.22** (0.08)
<i>high loan-to-value ratio</i> ($\text{loan-value ratio} \geq 0.95$)	0.189** (0.050)	1.49** (0.32)	0.79** (0.18)	0.79** (0.18)	0.84** (0.18)	0.79** (0.18)
<i>consumer credit score</i>	0.031** (0.005)	0.29** (0.04)	0.15** (0.02)	0.16** (0.02)	0.34** (0.11)	0.16** (0.02)
<i>mortgage credit score</i>	0.021 (0.011)	0.28* (0.14)	0.15* (0.07)	0.11 (0.08)	0.16 (0.10)	0.11 (0.08)
<i>public bad credit record</i>	0.197** (0.035)	1.23** (0.20)	0.70** (0.12)	0.70** (0.12)	0.72** (0.12)	0.70** (0.12)
<i>denied mortgage insurance</i>	0.702** (0.045)	4.55** (0.57)	2.56** (0.30)	2.59** (0.29)	2.59** (0.30)	2.59** (0.29)

Table 11.2, ctd.

<i>self-employed</i>	0.060** (0.021)	0.67** (0.21)	0.36** (0.11)	0.35** (0.11)	0.34** (0.11)	0.35** (0.11)
<i>single</i>				0.23** (0.08)	0.23** (0.08)	0.23** (0.08)
<i>high school diploma</i>				-0.61** (0.23)	-0.60* (0.24)	-0.62** (0.23)
<i>unemployment rate</i>				0.03 (0.02)	0.03 (0.02)	0.03 (0.02)
<i>condominium</i>					-0.05 (0.09)	
<i>black × P/I ratio</i>						-0.58 (1.47)
<i>black × housing expense-to-income ratio</i>						1.23 (1.69)
<i>Additional credit rating indicator variables</i>	no	no	no	no	yes	no
<i>constant</i>	-0.183** (0.028)	-5.71** (0.48)	-3.04** (0.23)	-2.57** (0.34)	-2.90** (0.39)	-2.54** (0.35)

(Table 11.2 continued)

Table 11.2, ctd.

(Table 11.2 continued)

F-Statistics and p-Values Testing Exclusion of Groups of Variables

	(1)	(2)	(3)	(4)	(5)	(6)
<i>Applicant single; HS diploma; industry unemployment rate</i>				5.85 (< 0.001)	5.22 (0.001)	5.79 (< 0.001)
<i>Additional credit rating indicator variables</i>					1.22 (0.291)	
<i>Race interactions and black</i>						4.96 (0.002)
<i>Race interactions only</i>						0.27 (0.766)
<i>Difference in predicted probability of denial, white vs. black (percentage points)</i>	8.4%	6.0%	7.1%	6.6%	6.3%	6.5%

These regressions were estimated using the $n = 2380$ observations in the Boston HMDA data set described in Appendix 11.1. The linear probability model was estimated by OLS, and probit and logit regressions were estimated by maximum likelihood. Standard errors are given in parentheses under the coefficients and p -values are given in parentheses under the F -statistics. The change in predicted probability in the final row was computed for a hypothetical applicant whose values of the regressors, other than race, equal the sample mean. Individual coefficients are statistically significant at the *5% or **1% level.

Ordered Probit: Course Evaluations and Beauty

We have the original continuous Y data (course evaluations) so we don't need to use these methods, but to illustrate ordered probit we construct artificially categorized data.

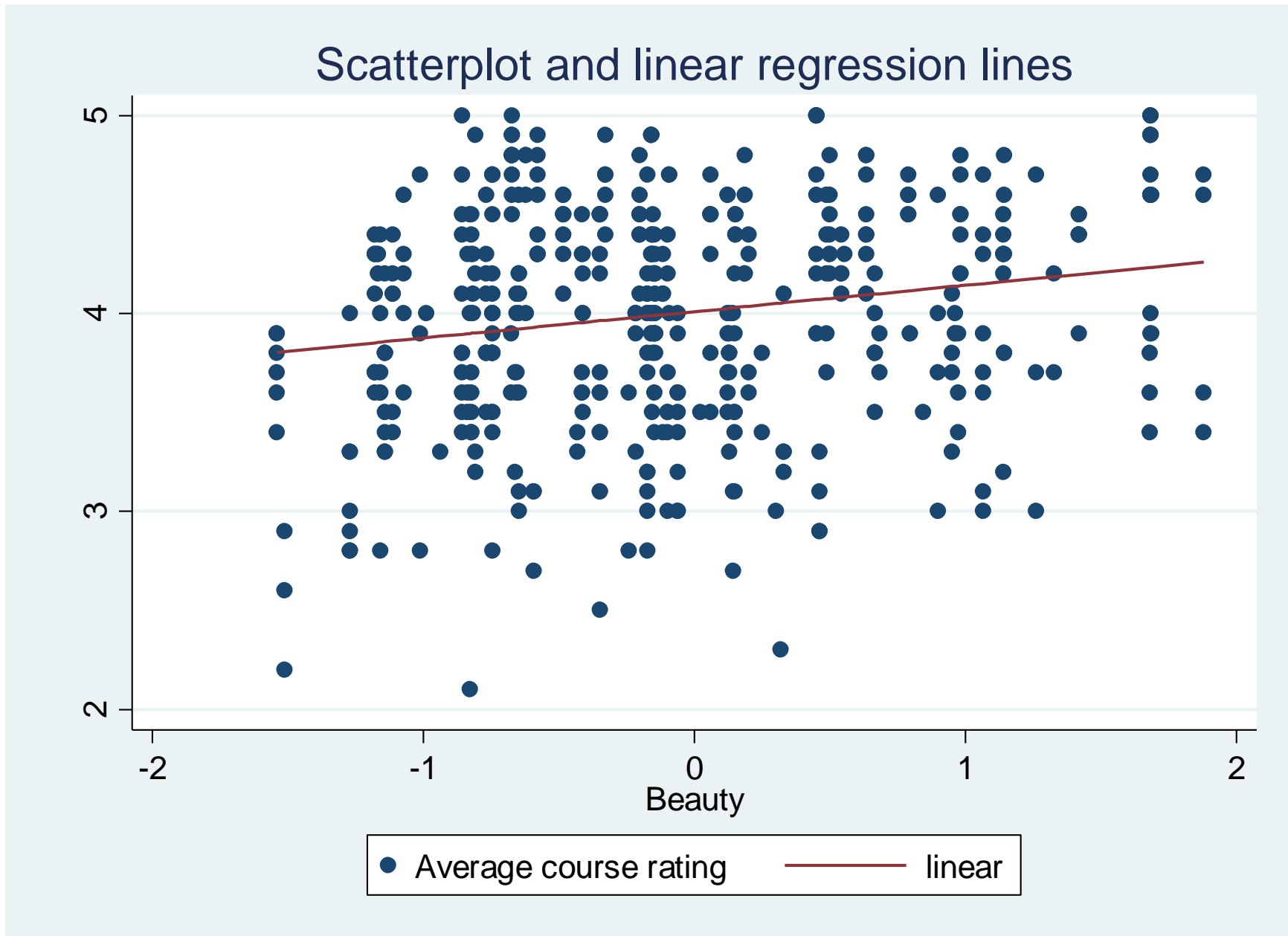
Artificial binary variable

$$eval_q234 = \begin{cases} 0 & \text{if } courseevaluation \text{ is in first quartile} \\ 1 & \text{if } courseevaluation \text{ is in top three quartiles} \end{cases}$$

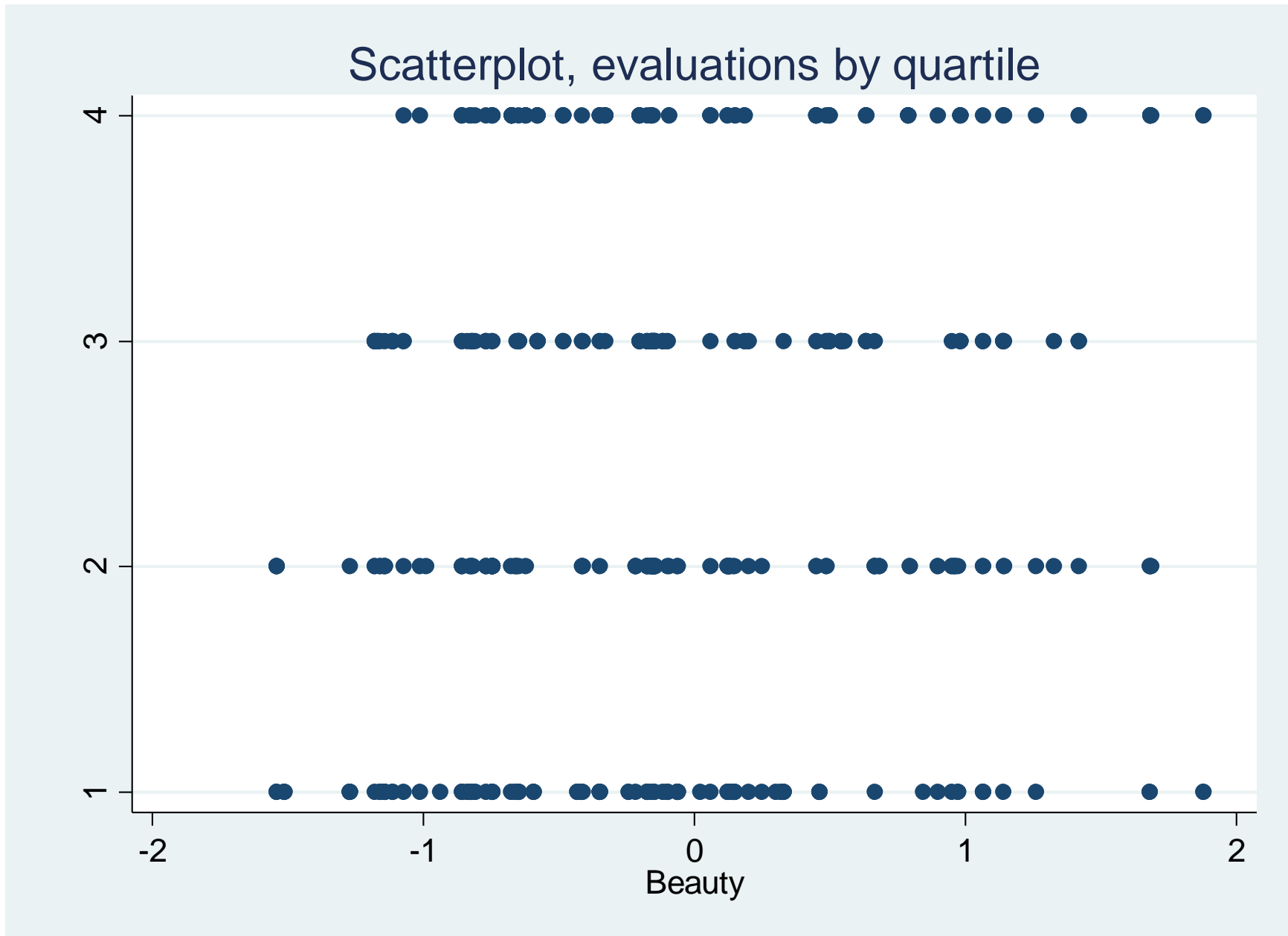
Artificial ordered categorical data

$$eval_ord = \begin{cases} 1 & \text{if } courseevaluation \text{ is in first quartile} \\ 2 & \text{if } courseevaluation \text{ is in second quartile} \\ 3 & \text{if } courseevaluation \text{ is in third quartile} \\ 4 & \text{if } courseevaluation \text{ is in fourth quartile} \end{cases}$$

Original data with linear regression:



Categorical course evaluation data (categorized by quartile)



STATA implementation – create variables; probit; ordered probit

```
. su courseevaluation, d;
```

Average course rating

```
-----
```

Percentiles		Smallest		
1%	2.6	2.1		
5%	3	2.2		
10%	3.3	2.3	Obs	463
25%	3.6	2.5	Sum of Wgt.	463
			Mean	3.998272
50%	4		Std. Dev.	.5548656
			Variance	.3078758
75%	4.4	5	Skewness	-.4658753
90%	4.7	5	Kurtosis	2.881628
95%	4.8	5		
99%	5	5		

```
. gen evalq2 = (courseevaluation>r(p25)) * (courseevaluation<=r(p50));
```

```
. gen evalq3 = (courseevaluation>r(p50)) * (courseevaluation<=r(p75));
```

```
. gen evalq4 = (courseevaluation>r(p75));
```

```
. gen eval_q234 = evalq2 + evalq3 + evalq4;
```

```
. gen eval_ord = 1 + evalq2 + 2*evalq3 + 3*evalq4;
```

```
. reg courseevaluation btystdave, r;
```

```
Linear regression
```

```
Number of obs = 463
F( 1, 461) = 16.94
Prob > F = 0.0000
R-squared = 0.0357
Root MSE = .54545
```

	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]	
courseeval~n						
btystdave	.1330014	.0323189	4.12	0.000	.0694908	.1965121
_cons	4.010023	.0253299	158.31	0.000	3.960246	4.059799

```
. reg eval_q234 btystdave, r;
```

```
Linear regression
```

```
Number of obs = 463
F( 1, 461) = 9.51
Prob > F = 0.0022
R-squared = 0.0194
Root MSE = .43833
```

	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]	
eval_q234						
btystdave	.078026	.0253052	3.08	0.002	.0282982	.1277538
_cons	.7412348	.0201643	36.76	0.000	.7016095	.7808601

```
. probit eval_q234 btystdave, r;
```

```
Iteration 0: log pseudolikelihood = -268.02744  
Iteration 1: log pseudolikelihood = -263.43691  
Iteration 2: log pseudolikelihood = -263.42781  
Iteration 3: log pseudolikelihood = -263.42781
```

Probit regression

```
Number of obs = 463  
Wald chi2(1) = 8.52  
Prob > chi2 = 0.0035  
Pseudo R2 = 0.0172
```

Log pseudolikelihood = -263.42781

	Coef.	Robust Std. Err.	z	P> z	[95% Conf. Interval]	
eval_q234						
btystdave	.2471247	.0846581	2.92	0.004	.081198	.4130515
_cons	.6597471	.0647791	10.18	0.000	.5327825	.7867117

```
. * ordered probit;
. oprobit eval_ord btystdave, r;
```

```
Iteration 0: log pseudolikelihood = -641.41106
Iteration 1: log pseudolikelihood = -633.59498
Iteration 2: log pseudolikelihood = -633.59449
```

Ordered probit regression

```
Number of obs = 463
Wald chi2(1) = 15.19
Prob > chi2 = 0.0001
Pseudo R2 = 0.0122
```

Log pseudolikelihood = -633.59449

```
-----+-----
```

eval_ord	Coef.	Robust Std. Err.	z	P> z	[95% Conf. Interval]	
btystdave	.2549661	.0654143	3.90	0.000	.1267564	.3831759
/cut1	-.6604092	.0638122			-.7854789	-.5353394
/cut2	.0227324	.0594761			-.0938386	.1393034
/cut3	.7111037	.0644798			.5847256	.8374819

```
-----+-----
```

Calculation of effects – ordered probit

Predicted probabilities for ordered probit (4 categories):

$$\Pr[Y_i = 0|X_i] = \Phi[c_1 - \beta_1 X_i]$$

$$\Pr[Y_i = 1|X_i] = \Phi[c_2 - \beta_1 X_i] - \Phi[c_1 - \beta_1 X_i]$$

$$\Pr[Y_i = 2|X_i] = \Phi[c_3 - \beta_1 X_i] - \Phi[c_2 - \beta_1 X_i]$$

$$\Pr[Y_i = 3|X_i] = 1 - \Phi[c_3 - \beta_1 X_i]$$

What is effect of increasing *btystdave* from -1 to 0 on probability of being in category 3?

$$\begin{aligned}x = -1: \quad PR[Y_i = 2|X_i = -1] &= \Phi[\hat{c}_3 - \hat{\beta}_1 \times (-1)] - \Phi[\hat{c}_2 - \hat{\beta}_1 \times (-1)] \\ &= \Phi[.711 - \mathbf{.255} \times (-1)] - \Phi[.023 - \mathbf{.255} \times (-1)] \\ &= \Phi[.966] - \Phi[.278] \\ &= .833 - .609 = .224\end{aligned}$$

$$\begin{aligned}x = 0: \quad PR[Y_i = 2|X_i = 0] &= \Phi[\hat{c}_3 - \hat{\beta}_1 \times 0] - \Phi[\hat{c}_2 - \hat{\beta}_1 \times 0] \\ &= \Phi[.711 - .255 \times 0] - \Phi[.023 - .255 \times 0] \\ &= \Phi[.711] - \Phi[.023] \\ &= .761 - .509 = .252\end{aligned}$$

An increase in *btystdave* from -1.0 to 0 is associated with an increase in the probability of being in the third quartile from .224 to .252, an increase of .028 percentage points

STATA .do file for Beauty example (probit, logit, ordered probit)

```
clear
capture log close
*****
* beauty_3_lect9.do
* Ec1123
* probit, ordered probit, illustrations
*****
set more off
log using beauty_3_oprobit_exs.log, replace
*****
* read in data
use hamermesh_beauty
desc
su
*
gen male = 1-female
gen bty2 = btystdave*btystdave
gen bty3 = btystdave*btystdave*btystdave
gen bty_male = btystdave*male
*
* create data for ordered probit - quartiles
su courseevaluation, d
gen evalq2 = (courseevaluation>r(p25)) * (courseevaluation<=r(p50))
gen evalq3 = (courseevaluation>r(p50)) * (courseevaluation<=r(p75))
gen evalq4 = (courseevaluation>r(p75))
gen eval_q234 = evalq2 + evalq3 + evalq4
```

```

gen eval_ord = 1 + evalq2 + 2*evalq3 + 3*evalq4
*
list courseevaluation eval_q234 eval_ord
*****
*   graphs
*****
reg courseevaluation btystdave, r
  predict peval
  label var peval "linear"
twoway scatter courseevaluation peval btystdave, ///
  ms(0 i i i) connect(. 1 1 1) sort(btystdave) ///
  title("Scatterplot and linear regression lines") ///
  xtitle("Beauty") ytitle("Course Overall") yscale(r(2 5))
graph export "beauty_3a.png", replace
*****
*   probit, logit regressions - one regressor
*****
reg courseevaluation btystdave, r
* linear probability model
reg eval_q234 btystdave, r
* probit
probit eval_q234 btystdave, r
* logit
logit eval_q234 btystdave, r
*
*****
*   ordered probit regressions - one regressor
*****
* ordered probit
oprobit eval_ord btystdave, r

```

```

twoway scatter eval_ord btystdave, ///
  ms(0 i i i) connect(. 1 1 1) sort(btystdave) ///
  title("Scatterplot, evaluations by quartile") ///
  xtitle("Beauty") ytitle("Course Overall")
graph export "beauty_3b.png", replace
sca a2 = _b[/cut2] - _b[btystdave]*(-1)
sca a3 = _b[/cut3] - _b[btystdave]*(-1)
dis a2 a3 normprob(a2) normprob(a3) normprob(a3)-normprob(a2)
sca b2 = _b[/cut2] - _b[btystdave]*(0)
sca b3 = _b[/cut3] - _b[btystdave]*(0)
dis b2 b3 normprob(b2) normprob(b3) normprob(b3)-normprob(b2)
*****
log close
clear
exit

```